Discrimination between Alzheimer Dementia and Controls by Automated Analysis of Multicenter FDG PET

K. Herholz,*¹ E. Salmon,† D. Perani,‡ J-C. Baron,§ V. Holthoff,|| L. Frölich,¶ P. Schönknecht,** K. Ito,††
R. Mielke,* E. Kalbe,* G. Zündorf,* X. Delbeuck,† O. Pelati,‡ D. Anchisi,‡ F. Fazio,‡ N. Kerrouche,§
B. Desgranges,§ F. Eustache,§ B. Beuthien-Baumann,‡‡ C. Menzel,§§ J. Schröder,** T. Kato,††
Y. Arahata,†† M. Henze,||| and W-D. Heiss*

* University Cologne, Neurological Clinic and Max-Planck-Institute for Neurological Research, 50931 Köln, Germany; † Cyclotron Research Centre and Service of Neurology, University of Liège, Belgium; ‡Institute of Neuroscience and Bioimaging-CNR, Institute H San Raffaele, University Vita-Salute HSR, University of Milano-Bicocca, Milan, Italy; §INSERM Unit 320, Cyceron, Caen, France; ||Department of Psychiatry and Psychotherapy, University of Technology, Dresden, Germany; ¶Department of Psychiatry and Psychotherapy, University Frankfurt am Main, Germany; ** Section of Geriatric Psychiatry, University of Heidelberg, Germany; ††National Institute of Longevity Sciences, Obu, Japan; ‡‡Department of Nuclear Medicine, University of Technology, Dresden, and PET-Center, Rossendorf, Germany; §§Department of Nuclear Medicine, University Frankfurt am Main, Germany; and ||||University of Heidelberg and German Cancer Research Center, Heidelberg, Germany

Received September 20, 2001

A new diagnostic indicator of FDG PET scan abnormality, based on age-adjusted t statistics and an automated voxel-based procedure, is presented and validated in a large data set comprising 110 normal controls and 395 patients with probable Alzheimer's disease (AD) that were studied in eight participating centers. The effect of differences in spatial resolution of PET scanners was minimized effectively by filtering and masking. In controls FDG uptake declined significantly with age in anterior cingulate and frontolateral perisylvian cortex. In patients with probable AD decline of FDG uptake in posterior cingulate, temporoparietal, and prefrontal association cortex was related to dementia severity. These effects were clearly distinct from age effects in controls, suggesting that the disease process of AD is not related to normal aging. Women with probable AD had significantly more frontal metabolic impairment than men. The new indicator of metabolic abnormality in AD-related regions provided 93% sensitivity and specificity for distinction of mild to moderate probable AD from normals, and 84% sensitivity at 93% specificity for detection of very mild probable AD (defined by Mini Mental Score 24 or better). All regions related to AD severity were already affected in very mild AD, suggesting that all vulnerable areas are affected to a similar degree already at disease onset. Ventromedial frontal cortex was also abnormal. In conclusion, automated analysis of multicenter FDG PET is feasible, provides insights

¹ To whom correspondence should be addressed at Max-Planck-Institut für neurologische Forschung, Gleueler Str. 50, 50931 Köln, Germany. Fax: +49-221-4726-298. E-mail: karl.herholz@pet.mpinkoeln.mpg.de. into AD pathophysiology, and can be used potentially as a sensitive biomarker for early AD diagnosis. • 2002 Elsevier Science (USA)

INTRODUCTION

The clinical diagnosis of Alzheimer's disease (AD) is defined by progressive memory impairment and other cognitive dysfunctions that are identified by medical and neuropsychological assessment. A major difficulty especially for large multicenter studies is the fact that neuropsychological instruments are often language dependent and culture specific, and single tests are mostly not comparable between different countries. Test results may also be influenced by the premorbid level of cognitive function. So far, no universal test battery for dementia is available except short screening tests, such as the Mini Mental Status Examination (MMSE) (Folstein et al., 1975), which must be complemented by additional tests for the diagnosis of dementia and has been criticized for a lack of sensitivity in early stages (Mungas et al., 1996; Grigoletto et al., 1999). Thus a biological indicator of AD that is not language and culture dependent continues to be of interest to facilitate objective diagnosis of AD.

Many studies over the past two decades have demonstrated that AD is associated with regional reductions of the cerebral metabolic rate of glucose (CMR-Glc) mainly in precuneus, temporo-parietal, and frontal association cortex (see reviews by Rapoport, 1991; Herholz, 1995; Perani and Cappa, 2001). FDG PET is very sensitive to detect CMRGlc reductions very early in AD, including in subjects at high genetic risk



but without clinical symptoms of dementia (Reiman et al., 1996; Small et al., 2000; de Leon et al., 2001). It is also more reliable for diagnostic purposes than single photon emission computed tomography (Mielke et al., 1994; Messa et al., 1994). The anatomical distribution of functional changes is closely related to impairment of some localized specific cognitive functions, such as aphasia, and their severity increases with progression of the disease (Haxby et al., 1990). CMRGlc in cerebellum, striatum, primary visual, and sensorimotor cortex is relatively spared, and therefore the typical impairment of associative cortices can be identified without absolute quantitation of CMRGlc. Thus, noninvasive assessment of FDG distribution without blood sampling usually is sufficient, and FDG PET at a resting state is easy to acquire and well tolerated by patients. Scanners and image reconstruction algorithms including correction for scatter and attenuation are technically mature, and their availability has increased considerably in many countries in recent years due to their increasing clinical use for detection of malignant tumors. Yet, at present, brain PET scans are often interpreted in a qualitative manner only by visual reading, which heavily depends on observer experience and training and lacks a clearly defined cutoff to distinguish between normal and pathological findings. Thus, FDG PET seems to be well suited as a biological marker for AD if complemented by an objective image analysis procedure that can be widely and easily applied in different PET centers.

Most previous FDG PET studies were conducted within single laboratories. Yet, in a previous multicenter study Herholz et al. (1993) demonstrated that it is possible to pool data derived from standard regions of interest. In that study, normal controls and patients with probable AD of mild to moderate severity were discriminated with high accuracy. In a subsequent study, deterioration in patients with mild cognitive impairment was predicted by metabolic abnormalities (Herholz et al., 1999). With the progress of voxel-based methods for anatomical image normalization and statistical parametric mapping (Minoshima et al., 1995; Friston et al., 1995a; Signorini et al., 1999), completely automated techniques for image analysis can be developed and applied to large multicenter data sets. We collected the largest sample of FDG PET scans in normals and patients with probable AD that has been described as yet from eight centers in five countries with four different languages. We report the development of a completely user-independent analysis and its results.

METHODS AND SUBJECTS

The study comprises 110 normal controls (70 male, 40 female) and 395 patients (169 male, 226 female) with probable AD according to NINCDS-ADRDA crite-

ria. Patients and controls were recruited at eight participating centers, who had performed clinical assessment, including neuropsychological testing and exclusion of other diseases by CT or MRI as required by NINCDS-ADRDA criteria, and PET scans in accordance with the study requirements. Neuropsychological testing batteries differed among centers, but it always included the Mini Mental Status Examination (Folstein et al., 1975) and, especially if MMSE was in the nondiagnostic range of 24 or higher, appropriate tests for immediate and delayed verbal memory, such as Buschke's selective reminding test (Grober et al., 1988) or the Wechsler memory scale (Wechsler, 1987), and verbal fluency were performed. In most patients testing was even more extensive including also specific tests for spatial memory, working memory, attention, language, and praxia and standard dementia rating scales (global deterioration scale (Reisberg et al., 1982), clinical dementia rating scale (Morris, 1993), or 2the Mattis dementia rating scale (Mattis, 1988; Fama et al., 1997)). Exclusion criteria were previous subarachnoid or intracerebral hemorrhage, intracranial tumors, hydrocephalus, all psychoses (including major depression), alcoholism, epilepsy, ischemic strokes, vascular dementia and other dementing illnesses, anemia, untreated thyroid dysfunction, renal insufficiency, and nonstabilized diabetes mellitus. Mean age of patients at the time of the PET study was 69 ± 8 years. Mean age at onset in patients was 66 ± 9 years (data not available in 64 patients), and mean MMSE score was 19 ± 6.

Healthy controls were recruited either by advertisement or as relatives of patients. They were interviewed to exclude dementia and all diseases that were exclusion criteria in the patient group. Most participating centers accepted treated medical conditions that are frequently present in clinically healthy controls such as arterial hypertension, hyperlipidemia, or compensated hypothyroidism. Controls who had subjective memory complaints were included if they had normal immediate and delayed memory tests, MMSE of 26 or higher, and no clinical evidence of dementia. Mean age of controls (57 \pm 13 years) was lower than that in patients; therefore an age-matched subset was selected for comparison with patients with probable AD.

Cases were acquired at eight centers, with PET scanners that differed with respect to field of view and spatial resolution (Table 1). Studies were done at a resting state with eyes closed and ears unplugged after intravenous injection of 110 to 370 MBq 18F-2-fluoro-2-deoxy-D-glucose (FDG). Images of tracer distribution in brain were used for analysis that had been summed up from part of a dynamic series of scans at most participating centers. Required minimum time interval between injection and scan start was 30 min. On average, scans were started 42 ± 19 min after injection. Required minimum scan duration was 10 min. On

TABLE 1

Participant ID	Scanner type (manufacturer)	Spatial resol	ution (mm)	Axial field of view (mm)	Number of cases	
		In-plane FWHM	Slice thickness		Controls	AD
1	ECAT EXACT (1)	6	3.375	162	15	51
	ECAT EXACT HR (1)	3.6	3.125	150	18	72
2	ADVANCE (2)	4.6	5.3	154	10	18
3	ECAT 951 (1)	8.5	3.37	108	25	188
4	ECAT EXACT HR+ (1)	3.6	2.46	155	12	27
6	ECAT EXACT HR+ (1)	3.6	2.46	155	7	21
7	ECAT EXACT 47 (1)	6.7	3.3	164	5	10
8	ECAT EXACT HR+ (1)	3.6	2.46	155	0	13
10	ECAT EXACT HR (1)	3.6	3.125	150	13	0

Overview on Participating Centers, Equipment, and Included Cases

Note. Manufacturers: (1) Siemens/CTI, (2) General Electric.

average, scan duration was 19 ± 3 min. Images were reconstructed using filtered backprojection including correction for attenuation (measured by transmission scan) and scatter using standard software as supplied by scanner manufacturers.

Basic image processing was done by MATLAB, using SPM99 routines (Wellcome Department of Cognitive Neurology, London, UK) for spatial normalization. Further image analysis was performed using newly developed IDL (Research Systems Incorporated, Boulder, CO) programs. MPITool (Pietrzyk *et al.*, 1996) was used for image display. SAS (SAS Institute, Inc., Cary, NC) was used for statistical evaluation of results.

All data were spatially normalized by affine 12-parameter transformation (Friston *et al.*, 1995a) using the SPM99 standard PET brain template. Normalized images were represented on a $79 \times 95 \times 68$ matrix with $2 \times 2 \times 2$ -mm voxel size. Images were then smoothed by a Gaussian filter of 12 mm FWHM that was limited to measured brain tissue using the IDL fast Fourier transform routine. To achieve that, the same filter was applied to a binary brain mask (with values 1 for measured voxels, 0 for nonmeasured voxels). The filtered image voxels were divided by the filtered binary mask values to compensate for signal loss due to the contribution of nonmeasured voxels (with zero signal). Thus, signal distortion due to filtering at the edges of measured data was avoided.

Special precautions were taken to minimize effects caused by differences of scanner resolution on the results. Smoothed normal scans obtained from high-resolution scanners (with axial resolution better than 6.0 mm) were compared to those from low-resolution scanners. A gray matter mask was defined by those voxels that had an intensity above whole-brain average on both scanner types. The mask contained 144,823 voxels (1159 ml). It did not include voxels from the uppermost 10 slices (i.e., from the top 20 mm of the brain) because confounding effects of spatial resolution were most prominent in these parts of the cortex. All images were scaled to an average intensity of 1 within this mask. In some instances there were scans that did not fill the entire mask due to unfortunate positioning of the head in the scanner. On average, 99.3% of the mask volume was within the field-of-view, with a minimum of 90.4% in one instance. Nonmeasured voxels were excluded from analysis, and statistical calculations were adjusted to a reduced number of subjects in voxels that were not present in all data sets. Intensity scaling of all images was also done after exclusion of nonmeasured voxels.

The influence of age and dementia severity (MMSE) was studied by regression analysis. The regression line and the standard error of the regression slope were determined. Student *t* values were calculated to test the significance of the regression (Steel and Torrie, 1980). Comparisons of FDG uptake between subject groups were also done on *t* maps. Local critical *t* values were calculated for a significance level of P = 0.05 (one-sided) with Bonferroni correction for multiple testing on 670 resolution elements (containing 216 voxels, as defined by the filter volume) within the mask. Only spatially contiguous clusters containing the volume of at least one resolution element (216 voxels) above that threshold were regarded as significant.

With the view of developing a potential diagnostic test we first established the normal range of FDG uptake in individual voxels. For that purpose, a random reference sample of 49 subjects was drawn from the whole sample of control subjects. Age regression was performed and the variance of predicted values was calculated for each voxel. Abnormal voxels were defined in individual images as those voxels whose values were lower than 95% age-adjusted prediction limits (Steel and Torrie, 1980), and corresponding t maps with reference to the values expected by the regression were calculated.

The next step was the development of an indicator whether an individual brain scan was normal or abnormal. This was necessary because even in normal scans approx. 5% of voxels (i.e., 7241 of 144,823) could be expected to show "abnormal" values lower than the 95% age-adjusted prediction limit. The sum of *t* values over all voxels with FDG uptake below the 95% ageadjusted prediction limit was selected as a global indicator of scan abnormality (*t* sum). To focus the search for abnormal voxels to those areas that are typical for Alzheimer's disease, a mask was defined by all voxels that had shown a close correlation (at P < 0.01 uncorrected) with MMSE in patients with probable AD (see Fig. 1). The sum over all *t* values of voxels with FDG uptake below the 95% age-adjusted prediction limit within this AD mask (AD *t* sum) was calculated for each individual.

In a third step, the normal range of these two indicators of scan abnormality, *t* sum and AD *t* sum, was studied. This was done in an independent sample comprising the remaining 61 normal subjects in order to avoid any bias that could have resulted from applying a criterion of abnormality to the same sample from which it had been derived from. Thus, the distribution of *t* sum and AD *t* sum was determined in subjects of the normal sample, who had not been used in the definition of abnormal voxels. Normal distribution was achieved by square-root transformation of *t* sum and AD *t* sum values prior to calculation of statistical tests.

Diagnostic accuracy for AD was determined by receiver operating characteristic (ROC) analysis (Metz, 1978). The indicators of scan abnormality, t sum and AD t sum, were tested for their ability to discriminate between patients with probable AD and 28 agematched controls, again including only controls from the second sample that had not been used for definition of abnormal FDG uptake.

RESULTS

Normal Controls

Average voxel values reflected the well-known distribution of CMRGlc. Local standard deviation showed some regional variations with an average of 0.09 across the whole masked brain, and values in the range 0.06 to 0.10 in most parts of cortex. Higher standard deviations up to a maximum of 0.26 were found in those parts of the basal ganglia that are adjacent to the ventricles, probably reflecting the high contrast and the spatial variability due to varying degrees of atrophy at the CSF-brain border that was not eliminated by spatial normalization. Standard deviation was also high with values of 0.15 to 0.20 at the occipital pole, the gray/white matter interface in frontal and parietal lobes, the middle part of cingulate gyrus, and in orbitofrontal cortex. These areas are known to show more variability of individual gyration that cannot completely be accounted for by affine spatial normalization than the other parts of the brain (Thompson et al., 1996, 1998).

There was a significant effect of age on local FDG uptake (Fig. 2). A significant decline with age was found in most parts of frontal cortex, most strongly in the anterior cingulate and in frontolateral perisylvian cortex bilaterally (Table 2a). A pronounced decline was also present in those parts of the basal ganglia that are adjacent to the ventricles, which are subject to incomplete recovery of activity because of atrophy effects (Mazziotta *et al.*, 1981) and had shown the high standard deviation in the analysis of average uptake. A relative increase with age, reflecting preservation of CMRGlc, was present in most parts of occipital, temporal, and sensorimotor cortex which was strongest in primary visual cortex.

Patients with Probable AD

Average FDG uptake in patients with probable AD demonstrated the typical distribution with temporoparietal and frontal reductions. Local standard deviation with an average of 0.14 across the masked brain was generally larger than that in controls, probably due to changes of FDG uptake caused by the disease.

The effect of dementia severity on FDG uptake (Fig. 2) was largest in an extended region that included temporoparietal cortex, posterior cingulate, and precuneus, and in frontal association cortex bilaterally (Table 2b). Some decline of FDG uptake was also noted in those parts of the basal ganglia that are adjacent to the ventricles, similar to the effect of age and probably related to atrophy. There was very little overlap between the brain regions with dementia-associated metabolic impairment and those with age-related metabolic decline in normals (red areas in Fig. 2: 392 voxels, representing 2% of all dementia-associated voxels).

Relative preservation of FDG uptake in dementia, resulting in a relative increase with MMSE, was observed in midbrain, putamen, insula, and sensorimotor and visual cortex (Table 2c). Previous studies that evaluated CMRGlc in absolute units had shown that there is no actual increase of CMRGlc in those regions but only relative preservation of CMRGlc (Herholz *et al.*, 1990). Thus, we used those regions (defined by P < 0.01, uncorrected) as a new reference to scale scans, instead of the global average, for the analyses described below.

A significant decline of FDG uptake in all brain regions that are typically affected in AD was already present in patients with probable AD with a MMSE score of 24 or higher (Figs. 3 and 4). A significant reduction of FDG uptake was also noted in a small region in ventromedial frontal cortex (Table 2d). At more severe stages of the dementia (i.e., MMSE <24), impairment in these areas became significantly more severe, but the spatial distribution of changes remained essentially unchanged.



FIG. 1. Normal brain FDG uptake in shown in a gray scale within the mask used for data analysis. Regions that show decline of FDG uptake correlated with dementia severity in patients with probable AD are marked by red overlay; regions that have relatively preserved FDG uptake in AD are marked by blue overlay. Slices are 2 mm thick and parallel to the AC-PC level which is marked by a box.



TABLE 2

	0						
<u> </u>		Stereotactic coord. of maximum					
Size (voxels)	Maximum t	x	У	Z	Anatomical structures		
		(a) Ag	ge-related decline	e of FDG uptake	e in controls		
6998	10.0198	-2	26	36	Medial and superior frontal gyrus, right and left		
1231	8.52082	-40	18	2	Left inferior frontal gyrus		
1146	8.41626	-8	6	18	Left head of caudate/ventricle interface		
1448	8.18099	44	18	-4	Right inferior frontal gyrus		
294	7.02169	12	6	18	Right head of caudate/ventricle interface		
313	5.68729	28	28	60	Right superior frontal gyrus		
		(b) MMSE-relate	ed decline of FDC	G uptake in pati	ents with probable AD		
18623	9.65401	-64	-42	-6	Left and right temporoparietal association cortex, posterior cingulate and precuneus		
1151	6.24284	-30	16	66	Left superior frontal gyrus		
850	5.94052	38	18	62	Right superior frontal gyrus		
	(c) Preserv	ed FDG uptake i	in patients with	probable AD (re	lative increase related to MMSE)		
3560	9.85866	38	-4	16	Right putamen, midbrain, insula, and motor cortex		
5313	9.70108	-22	-22	18	Left putamen, midbrain, insula, and motor cortex		
11156	9.07434	8	-60	-8	Left and right visual cortex and cerebellum		
2088	8.60933	-10	-8	68	Left and right superior medial frontal gyrus		
	(d) Reduced FDG	uptake in patien	ts with very mile	d AD (MMSE 24	4 or higher) compared to normal controls		
22554	9.58125	-4	-58	32	Right and left post. cingulate, precuneus and left temporoparietal and frontal association cortex		
6812	7.39844	52	-62	52	Right temporoparietal association cortex		
1985	6.20894	48	16	36	Right frontal association cortex		
316	4.97733	-2	32	-10	Right and left ventromedial frontal cortex		
					-		

Location of Significant Voxel Clusters (Thresholded at P < 0.05 with Correction for Multiple Comparisons)

Discrimination between Controls and Patients with Probable AD

First, possible confounding effects of age and scanner resolution were studied in controls. In whole brain t sum tended to be higher in high-resolution scanners than in low-resolution scanners (Table 3a), but differences were not significant and there was no effect of age. When analysis was limited to those regions that had shown a significant relation with dementia severity (AD t sum, Table 3b), these values tended to be higher in younger controls, but again differences were not significant, and there was no effect of scanner resolution. Mean control values were 15,419.6 \pm 9096.4 for t sum as a global measure of abnormality, and 4673.0 \pm 3187.25 for AD t sum as a measure of abnormality in AD-related regions.

Patients with probable AD had much higher t sum and AD t sum scores than normals (Table 4). There were no differences with respect to scanner resolution, but patients with early-onset dementia (age at onset < 65 years) had significantly higher t sum and AD t sum scores (P < 0.0001), even after adjustment for the slightly more severe dementia in the early-onset group (MMSE scores 16.6 ± 6.2 vs 18.2 ± 5.4 , P = 0.04). The *t*-score sums were significantly related to dementia severity (r = 0.56, P < 0.0001 for *t* sum, and r = 0.61, P < 0.0001 for AD *t* sum, Fig. 5). There was no significant effect of gender on *t*-score sums, although values tended to be slightly lower in women.

AD *t* sum provided the best discrimination between patients with probable AD and age-matched controls. The area under the ROC curve (AUC) was 0.97 (Fig. 6). At an AD *t* sum threshold of 10,150 sensitivity was 93% at a specificity of 93%. This threshold was very similar to the 95% confidence limit of AD *t*-sum distribution in normal controls at 10,953. AD *t* sums generally corresponded well with visual appearance of the scans. As illustrated in Fig. 7, scans with AD *t* sums of 10,000 or less clearly appeared normal, and those with AD *t* sums of 20,000 or higher clearly appeared abnormal, whereas in the range between 10,000 and 20,000 visual appearance would have been open for subjective interpretation. The two controls misclassified as AD by AD *t* sum had values of 10,768 and 11,240 just above the



FIG. 3. Difference (*t* map) between age-matched controls and very mild AD with MMSE scores of 24 or higher. The brightest areas mark the extent of significant changes with cross hairs at the local extreme values (see Table 2d).

critical threshold with very mild abnormalities in temporoparietal cortex. In all patients with probable AD but normal AD *t* sums visual examination of scans also did not reveal abnormalities that would have been judged as typical for AD. In four cases focal metabolic reductions of unknown clinical significance were present: left thalamus and frontal cortex, frontal cortex (associated with severe general atrophy), left parahippocampal gyrus, and left hippocampus and frontal cortex.

In patients with very mild dementia (MMSE 24 or higher) sensitivity was still 84% (at the same specificity, AUC 0.93). Even when only patients with MMSE of 27 or higher (average MMSE 27.8 \pm 0.7, age 65.5 \pm 6.6,

n = 17) were selected and matched to normal controls in whom MMSE scores had been recorded (average MMSE 28.1 ± 1.1, age 64.5 ± 8.0, n = 17), probable AD patients were identified with 83% sensitivity and 82% specificity (AUC 0.89). In senile patients and controls (age 65 or higher) discrimination was maintained at a high level (93% sensitivity at 90% specificity, AUC 0.97), in spite of the somewhat lower average AD *t*-sum scores in late-onset dementia. Discrimination by *t*-sum values over all abnormal voxels was generally less accurate with a sensitivity of 87% (AUC 0.95) in the whole patient group and 72% (AUC 0.90) in very mild dementia.



FIG. 4. *t* map of a patient with very mild AD (MMSE 27, age 75 years). Significant voxels (P < 0.05, uncorrected) are marked white, and maxima of connected clusters are marked by cross hairs. *t*-sum (26953.5) was within normal limits, but AD *t* sum (17369.6) clearly exceeded the age-adjusted 95% prediction limit of 11089.8.

Anatomical Distribution of Abnormalities in AD

We also checked the relative contribution of frontal, temporo-parietal, and posterior cingulate areas to the total AD *t* sum in patients (Table 5). More than 50% of metabolic reductions were located in temporoparietal regions. Women had significantly more frontal impairment (0.29 ± 0.11) than men (0.27 ± 0.12 , P = 0.02 in *t* test), whereas men had more temporo-parietal impairment (0.55 ± 0.10) than women (0.52 ± 0.10 , P = 0.01). There was no significant relation of the regional distribution with dementia severity or age at onset. Yet, metabolic impairment was significantly more se-

vere in left hemispheric regions (asymmetry index (l - r)/(l + r) was 0.03 ± 0.11 , P < 0.001 in *t* test). Regression of the asymmetry index on MMSE showed a tendency for more prominent left impairment in mild AD, but neither intercept nor slope was significant.

DISCUSSION

In the present study we developed and validated a new quantitative parameter, the sum over all *t* values of abnormal voxels in the AD mask (AD *t* sum), which provides a measure of the depth and extent of meta-

TABLE 3

Abnormal V	Voxel 1	Indices	in	Controls	by	Age a	and
	Sca	anner R	eso	olution			

		Resol	Resolution		
Age	t sum	High	Low	All	
((a) Sum of t so	cores across all a	bnormal voxels (t sum)	
≤60	N	20.00	13.00	33.00	
	Mean	16425.07	13941.30	15446.61	
	Std	7910.01	7474.12	7721.22	
>60	N	13.00	15.00	28.00	
	Mean	17676.74	13404.00	15387.77	
	Std	9411.00	11547.76	10640.34	
All	N	33.00	28.00	61.00	
	Mean	16918.15	13653.46	15419.60	
	Std	8411.20	9697.79	9096.39	
	(b) Sum of t s	cores in AD-relat	ed regions (AD t	sum)	
≤60	N	20.00	13.00	33.00	
	Mean	5232.02	4861.50	5086.06	
	Std	3057.53	3031.95	3005.29	
>60	N	13.00	15.00	28.00	
	Mean	4249.07	4131.73	4186.21	
	Std	2824.15	3895.31	3378.77	
All	N	33.00	28.00	61.00	
	Mean	4844.80	4470.55	4673.01	
	Std	2963.02	3477.17	3187.25	

bolic abnormality in brain regions that are typically affected in AD. It represents the severity of functional impairment of the association cortices, is significantly related to dementia severity, and discriminates between normals and patients with high accuracy. The approach combines the virtues of voxel-based parametric mapping with the diagnostic information that can be drawn from the topographic distribution of abnormal findings. It is an alternative to the use of ratios (Herholz *et al.*, 1990; Minoshima *et al.*, 1995) or multivariate procedures (Kippenhan *et al.*, 1992; Ichimiya *et al.*, 1994; Rapoport, 1997).

The present study confirms the high sensitivity of FDG PET to detect AD that has also been found in previous studies by visual rating (Salmon *et al.*, 1994), by a metabolic ratio based on regions of interest (Herholz et al., 1993), and by stereotactic surface projection analysis (Minoshima et al., 1995). In these studies, clinical diagnosis of probable AD was used as a standard which may not be accurate in all cases. Although low accuracy was reported in community-based case series (Mendez et al., 1992; Lim et al., 1999), reasonable accuracy in the range 87 to 100% is usually achieved when clinical diagnosis is established according to standard criteria in academic hospitals (Morris et al., 1988). High sensitivity (93%) of FDG PET was also confirmed in a series of 138 patients with pathological confirmation of AD (Silverman et al., 2001). High sensitivity for very early AD is evident from stud-

TABLE 4

Abnormal	Voxel	Indices	in l	Patients	by A	Age at	Onset	and
		Scann	er	Resolution	on			

		Resol	Resolution			
Age	t sum	High	Low	All		
	(a) Sum of t	scores across all	abnormal voxels	(<i>t</i> sum)		
<65	N	61.00	72.00	133.00		
	Mean	128119.43	120591.43	124044.12		
	Std	78711.54	71970.87	74942.81		
≥ 65	N	60.00	138.00	198.00		
	Mean	87242.96	73323.10	77541.24		
	Std	46627.43	45392.72	46099.87		
NA	N	30.00	34.00	64.00		
	Mean	85734.68	91447.20	88769.46		
	Std	77441.56	51056.93	64298.84		
All	N	151.00	244.00	395.00		
	Mean	103456.31	89796.62	95018.43		
	Std	70058.09	58860.96	63636.13		
	(b) Sum of t-	scores in AD-rela	ated regions (AD-	t-sum)		
<65	N	61.00	72.00	133.00		
	Mean	78868.44	74842.09	76688.76		
	Std	50156.91	45029.81	47309.73		
≥ 65	N	60.00	138.00	198.00		
	Mean	46912.29	41945.94	43450.89		
	Std	28296.76	29653.26	29266.79		
NA	N	30.00	34.00	64.00		
	Mean	40427.72	51813.90	46476.63		
	Std	40746.98	34028.66	37464.85		
All	N	151.00	244.00	395.00		
	Mean	58533.40	53028.04	55132.62		
	Std	43931.43	38162.74	40499.00		

ies of asymptomatic subjects at high genetic risk for AD, which showed FDG PET abnormalities in areas typical for AD (Reiman *et al.,* 1996; Small *et al.,* 2000; de Leon *et al.,* 2001).

Specificity is a more complex issue. With respect to the distinction from normal controls, values in the



FIG. 5. Regression of *t* score sum in AD-related areas vs dementia severity. Solid circles denote early-onset, open circles late-onset AD, and diamonds patients in whom age at onset was not available. The range of values found in age-matched controls is shaded.



false positives

FIG. 6. ROC curve for diagnostic discrimination between agematched controls (not including controls in whom normal range was determined) and patients with probable AD. The dashed curve represents the analysis in very mild AD (MMSE \geq 24).

TABLE 5

Proportion of Abnormal t values within AD-Typical Regions

	Mean	Std
Left frontal	0.15	0.07
Right frontal	0.13	0.06
Left temppar.	0.28	0.08
Right temppar.	0.25	0.09
Precuneus	0.19	0.07

same range as for sensitivity (90% or higher) were usually found (Herholz *et al.*, 1993; Minoshima *et al.*, 1995) in accordance with the present series. Yet, with the present data we cannot determine the specificity for distinction of AD from other types of dementia. The autopsy data that are available so far suggest that it may be in the range 63 to 78% (Hoffman *et al.*, 2000;



FIG. 7. Comparison of spatially normalized FDG uptake (top of each triplet, gray scale of full intensity range with 20% background subtraction, transaxial slice 42 mm above AC-PC plane), corresponding deviation from normal reference sample (maximum intensity voxels represent significant reduction at P < 0.05, uncorrected), and associated *t*-sum scores (values rounded) in a representative sample of 12 individual scans from the study. Scans within the box (right bottom) are from normal individuals; all others are from patients with probable AD.

Silverman *et al.*, 2001). Thus, in patients presenting with cognitive symptoms, regional brain metabolism is sensitive to the presence of AD, and to neurodegenerative disease generally. A negative PET scan helps provide assurance that cognitive abilities are likely to remain stable in the next 2 years following the study (Herholz *et al.*, 1999).

The consistency of the present study with previous findings suggests that the results of the present method are well comparable to other methods, such as visual scan interpretation by experienced observers (see also Fig. 7), ROI-based analysis (Herholz *et al.*, 1990), and standardized surface projections (Minoshima *et al.*, 1995). Yet, the present study demonstrates for the first time use of a completely automated method in a large multicenter study.

Metabolic reduction in ventromedial frontal cortex has not been noted previously with PET and its detection in this study is probably due to the statistical power provided by the large sample size. It may correspond to neurofibrillary tangles and Alz 50 immunoreactive neurons that have been found in this area (Chu et al., 1997). Otherwise, location of reduced FDG uptake in AD in this study was very similar to that found in previous studies (see Herholz, 1995, for a review of ROI data, and Minoshima et al., 1995, and Signorini et al., 1999, for voxel data). Yet, we could not confirm a preferential affection of the posterior cingulate and precuneus in early AD, but found that the relative proportion of abnormalities in posterior cingulate, temporo-parietal, and fronto-lateral association cortex was independent from disease severity, suggesting that, in a patient population, all of these vulnerable areas are affected right from the outset. Across patients, the metabolic impairment seems to be more narrowly and consistently localized in the precuneus and posterior cingulate than in the more extended frontal and temporoparietal association cortex, which could be the reason that maximum reduction in early AD is found there by parametric mapping (see Table 2d and Minoshima et al., 1995). The more widespread temporo-lateral and frontal impairment was demonstrated with SPM in smaller AD samples (Desgranges et al., 1998; Salmon et al., 2000) less consistently because it affects only the same region but not necessarily the same voxels in different patients.

Impairment of frontolateral association cortex was stronger in women than in men. To our knowledge, this finding has not been reported before with PET and its detection may also be due to the higher power of our large sample. It is currently unclear whether it is related to inferior performance of women with AD in naming and verbal fluency tasks and in delayed recall that has been noted in a large neuropsychological study (Henderson and Buckwalter, 1994), and further detailed comparisons between FDG PET and neuropsychological testing are required. Similar to previous studies (Loewenstein *et al.*, 1989), we found more severe abnormalities in the left than in the right hemisphere. This asymmetry tended to be more prominent in mild dementia. We therefore assume that it may be due to a selection bias caused by the fact that left hemispheric impairment is often related to aphasic disturbance (Haxby *et al.*, 1985), which may more frequently lead to referral for PET studies than the visuospatial deficits that are related to right hemispheric impairment and may more easily be tolerated in mild dementia.

The comparison of the locations of dementia-associated metabolic impairment and age-related decline in normals showed that they are quite distinct. This suggests that AD is a disease process that is not related to normal aging. In the frontal lobe, normal aging had its largest effect in frontomesial cortex, whereas the ADrelated changes were mostly in fronto-lateral cortex. Similar effects of normal aging on CMRGlc had been described before (Kuhl et al., 1984; Petit-Taboue et al., 1998; Garraux et al., 1999), but had not yet been compared directly to AD-related changes. Among patients, late onset of disease was associated with less prominent local abnormalities, which is in accordance with previous studies (Mielke et al., 1992; Salmon et al., 2000), but had little consequence in the accuracy of discrimination between controls and patients above age 65.

The present approach follows the concept of statistical parametric imaging (SPM) (Friston et al., 1995b), and actually uses the procedure provided by SPM99 for spatial normalization. Yet, it is the first study to apply this approach to a data set of several hundred subjects, which required new special program code (implemented using IDL) to handle the amount of data. In addition, we introduced t sum and AD t sum as new statistical indicators of global and AD-like scan abnormality. In contrast to comparisons of groups, which is the standard application of traditional statistical parametric mapping, these parameters provide a sensitive measure of individual scan abnormality which includes an adjustment for age-related changes (see Fig. 4 for an illustration). We validated them for future use as a clinical diagnostic instrument by using two separate subsets of normals, one for development of the indicators, and the other for testing them.

The AD *t* sum was calculated from those regions that had shown a significant relation with dementia severity. To exclude potential overestimation of diagnostic accuracy due to the fact that those regions had been determined from the same sample, we also identified the regions with significant relation to dementia severity in a random subset of 103 patients with probable AD and performed the ROC analysis of diagnostic discrimination only in the remaining patients. The result was almost identical with sensitivity 94% and specificity 93%.

We also solved the problem of different scanner resolution (Grady et al., 1989), which in the past has largely obviated data pooling and voxel-based multicenter PET studies. Of course, the effects of different scanner resolution could not be eliminated completely and were still recognizable as trends (see Table 3a), but in contrast to the huge significances that were obtained for the contrast between controls and AD, they were negligible. This was achieved by creation of a special mask that includes only the ridge of grey matter representation in the smoothed scans. Scanners with lower spatial resolution underestimate grav matter uptake due to contamination with surrounding low signal more than high-resolution scanners, and this effect is stronger in small structures than in larger ones (Mazziotta et al., 1981). Yet, with the additional heavy smoothing by a 12-mm Gaussian filter the effect becomes uniform enough in the gray matter ridge to result mainly in a globally lower signal which is then scaled to the same value in all subjects. As a result, we did not find significant resolution-dependent effects in normals and in the discrimination of patients from normals. There is certainly hope that with more widespread use of high-resolution scanners only one could even detect smaller abnormal volumes and thus improve diagnostic sensitivity, but that would also require further progress with spatial normalization which currently also requires smoothing to reduce residual signal variation caused by incomplete interindividual gyrus matching after affine transformation. Analysis of high-resolution data and less smoothing probably would also be required to detect metabolic abnormalities of hippocampal structures (Jagust et al., 1993).

Partial volume effects due to cortical atrophy also lead to underestimation of local cerebral glucose metabolism (Meltzer et al., 1999). Although we did not correct explicitly for that, the effects of age-related atrophy were taken into account by adjustment of the normal reference values by linear regression on age. Local atrophy due to disease, in this case in particular due to AD, does contribute to a reduction of the FDG PET signal in affected regions (Baron et al., 2001). Since we were not primarily interested in measuring accurate CMRGlc in this study, but to detect AD, it was not necessary or even desirable to correct for an effect that is disease-related and possibly may improve diagnostic sensitivity. Previous studies (Alavi et al., 1993; Meltzer et al., 1996; Ibanez et al., 1998) demonstrated that, even after correction for atrophy, the reduction of CMRGlc persists in temporal and cingulate cortex and thus is not merely a consequence of atrophy. Thus, early cortical functional impairment that may precede atrophy is detected with high sensitivity by FDG PET, and in later stages atrophy may additionally contribute to signal loss because of partial volume effects.

The study opens the perspective to use this reference data base and the completely user-independent analysis of FDG PET in other diagnostic, longitudinal, and intervention studies of Alzheimer's disease and related conditions. Since the present study was based on clinical diagnosis as the gold standard, further studies are required to clarify the additional value of FDG PET with this automatic procedure over clinical assessment in predicting clinical decline and the histopathological diagnosis of AD in patients who present to the clinic with early dementia. The procedure is available for further scientific evaluation by submission of FDG PET resting brain images via FTP to the NEST-DD coordinator (see project homepage at www.nestdd.org).

ACKNOWLEDGMENT

This study was conducted by the Network for Efficiency and Standardisation of Dementia Diagnosis (NEST-DD) with support from the European Commission (Framework V). Work at CRC Liège was also supported by grants from FNRS and FMRE, Belgium.

REFERENCES

- Alavi, A., Newberg, A. B., Souder, E., and Berlin, J. A. 1993. Quantitative analysis of PET and MRI data in normal aging and Alzheimer's disease: Atrophy weighted total brain metabolism and absolute whole brain metabolism as reliable discriminators. J. Nucl. Med. 34: 1681–1687.
- Baron, J. C., Chetelat, G., Desgranges, B., Perchey, G., Landeau, B., de l, S. V., and Eustache, F. 2001. In vivo mapping of gray matter loss with voxel-based morphometry in mild alzheimer's disease. *NeuroImage* 14: 298–309.
- Chu, C. C., Tranel, D., Damasio, A. R., and Van Hoesen, G. W. 1997. The autonomic-related cortex: Pathology in Alzheimer's disease. *Cereb. Cortex* **7**: 86–95.
- de Leon, M. J., Convit, A., Wolf, O. T., Tarshish, C. Y., DeSanti, S., Rusinek, H., Tsui, W., Kandil, E., Scherer, A. J., Roche, A., Imossi, A., Thorn, E., Bobinski, M., Caraos, C., Lesbre, P., Schlyer, D., Poirier, J., Reisberg, B., and Fowler, J. 2001. Prediction of cognitive decline in normal elderly subjects with 2-F-18-fluoro-2-deoxy-D-glucose positron-emission tomography (FDG PET). *Proc. Natl. Acad. Sci.USA* **98**: 10966–10971.
- Desgranges, B., Baron, J. C., de l, S., Petit-Taboue, M. C., Benali, K., Landeau, B., Lechevalier, B., and Eustache, F. 1998. The neural substrates of memory systems impairment in Alzheimer's disease. A PET study of resting brain glucose utilization. *Brain* **121**: 611– 631.
- Fama, R., Sullivan, E. V., Shear, P. K., Marsh, L., Yesavage, J. A., Tinklenberg, J. R., Lim, K. O., and Pfefferbaum, A. 1997. Selective cortical and hippocampal volume correlates of Mattis Dementia Rating Scale in Alzheimer disease. *Arch. Neurol.* 54: 719–728.
- Folstein, M. F., Folstein, S. E., and McHugh, P. R. 1975. "Minimental state." A practical method for grading the cognitive state of patients for the clinician. *J. Psychiatr. Res.* **12**: 189–198.
- Friston, K. J., Ashburner, J., Poline, J. B., Frith, C. D., Heather, J. D., and Frackowiak, R. S. J. 1995a. Spatial registration and normalisation of images. *Human Brain Mapping* 2: 165–189.
- Friston, K. J., Holmes, A. P., Worsley, K. J., Poline, J. B., Frith, C. D., and Frackowiak, R. S. J. 1995b. Statistical parametric maps in

functional imaging: a general linear approach. *Human Brain Mapping* **2:** 189–210.

- Garraux, G., Salmon, E., Degueldre, C., Lemaire, C., Laureys, S., and Franck, G. 1999. Comparison of impaired subcortico-frontal metabolic networks in normal aging, subcortico-frontal dementia, and cortical frontal dementia. *NeuroImage* **10**: 149–162.
- Grady, C. L., Berg, G., Carson, R. E., Daube-Witherspoon, M. E., Friedland, R. P., and Rapoport, S. I. 1989. Quantitative comparison of cerebral glucose metabolic rates from two positron emission tomographs. *J. Nucl. Med.* **30**: 1386–1392.
- Grigoletto, F., Zappala, G., Anderson, D. W., and Lebowitz, B. D. 1999. Norms for the Mini-Mental State Examination in a healthy population. *Neurology* **53**: 315–320.
- Grober, E., Buschke, H., Crystal, H., Bang, S., and Dresner, R. 1988. Screening for dementia by memory testing. *Neurology* **38**: 900–903.
- Haxby, J. V., Duara, R., Grady, C. L., Cutler, N. R., and Rapoport, S. I. 1985. Relations between neuropsychological and cerebral metabolic asymmetries in early Alzheimer's disease. *J. Cereb. Blood Flow Metab.* 5: 193–200.
- Haxby, J. V., Grady, C. L., Koss, E., Horwitz, B., Heston, L., Schapiro, M., Friedland, R. P., and Rapoport, S. I. 1990. Longitudinal study of cerebral metabolic asymmetries and associated neuropsychological patterns in early dementia of the Alzheimer type. *Arch. Neurol.* 47: 753–760.
- Henderson, V. W., and Buckwalter, J. G. 1994. Cognitive deficits of men and women with Alzheimer's disease. *Neurology* 44: 90–96.
- Herholz, K. 1995. FDG PET and differential diagnosis of dementia. Alzheimer Dis. Assoc. Disorders **9:** 6–16.
- Herholz, K., Adams, R., Kessler, J., Szelies, B., Grond, M., and Heiss, W. D. 1990. Criteria for the diagnosis of Alzheimer's disease with positron emission tomography. *Dementia* 1: 156–164.
- Herholz, K., Nordberg, A., Salmon, E., Perani, D., Kessler, J., Mielke, R., Halber, M., Jelic, V., Almkvist, O., Collette, F., Alberoni, M., Kennedy, A., Hasselbalch, S., Fazio, F., and Heiss, W.-D. 1999. Impairment of neocortical metabolism predicts progression in Alzheimer's disease. *Dementia Geriatr. Cog. Disord.* 10: 494–504.
- Herholz, K., Perani, D., Salmon, E., Franck, G., Fazio, F., Heiss, W. D., and Comar, D. 1993. Comparability of FDG PET studies in probable Alzheimer's disease. J. Nucl. Med. 34: 1460–1466.
- Hoffman, J. M., Welsh-Bohmer, K. A., Hanson, M., Crain, B., Hulette, C., Earl, N., and Coleman, R. E. 2000. FDG PET imaging in patients with pathologically verified dementia. *J. Nucl. Med.* **41**: 1920–1928.
- Ibanez, V., Pietrini, P., Alexander, G. E., Furey, M. L., Teichberg, D., Rajapakse, J. C., Rapoport, S. I., Schapiro, M. B., and Horwitz, B. 1998. Regional glucose metabolic abnormalities are not the result of atrophy in Alzheimer's disease. *Neurology* **50**: 1585–1593.
- Ichimiya, A., Herholz, K., Mielke, R., Kessler, J., Slansky, I., and Heiss, W. D. 1994. Difference of regional cerebral metabolic pattern between presenile and senile dementia of the Alzheimer type: A factor analytic study. *J. Neurol. Sci.* **123**: 11–17.
- Jagust, W. J., Eberling, J. L., Richardson, B. C., Reed, B. R., Baker, M. G., Nordahl, T. E., and Budinger, T. F. 1993. The cortical topography of temporal lobe hypometabolism in early Alzheimer's disease. *Brain Res.* 629: 189–198.
- Kippenhan, J. S., Barker, W. W., Pascal, S., Nagel, J., and Duara, R. 1992. Evaluation of a neural-network classifier for PET scans of normal and Alzheimer's disease subjects. *J. Nucl. Med.* 33: 1459– 1467.
- Kuhl, D. E., Metter, E. J., Riege, W. H., and Hawkins, R. A. 1984. The effect of normal aging on patterns of local cerebral glucose utilization. *Ann. Neurol.* 15(Suppl): S133–137.
- Lim, A., Tsuang, D., Kukull, W., Nochlin, D., Leverenz, J., Mccormick, W., Bowen, J., Teri, L., Thompson, J., Peskind, E. R., Ras-

kind, M., and Larson, E. B. 1999. Clinico-neuropathological correlation of Alzheimer's disease in a community-based case series. *J Am. Geriatr. Soc.* **47:** 564–569.

- Loewenstein, D. A., Barker, W. W., Chang, J. Y., Apicella, A., Yoshii, F., Kothari, P., Levin, B., and Duara, R. 1989. Predominant left hemisphere metabolic dysfunction in dementia. *Arch. Neurol.* 46: 146–152.
- Mattis, S. 1988 *Dementia Rating Scale (DRS).* Psychological Assessment Resources, Inc., Odessa, FL.
- Mazziotta, J. C., Phelps, M. E., Plummer, D., and Kuhl, D. E. 1981. Quantitation in positron emission computed tomography: 5. Physical–Anatomical effects. *J. Comput. Assist. Tomogr.* 5: 734–743
- Meltzer, C. C., Kinahan, P. E., Greer, P. J., Nichols, T. E., Comtat, C., Cantwell, M. N., Lin, M. P., and Price, J. C. 1999. Comparative evaluation of MR-based partial-volume correction schemes for PET. J. Nucl. Med. 40: 2053–2065.
- Meltzer, C. C., Zubieta, J. K., Brandt, J., Tune, L. E., Mayberg, H. S., and Frost, J. J. 1996. Regional hypometabolism in Alzheimer's disease as measured by positron emission tomography after correction for effects of partial volume averaging. *Neurology* **47**: 454– 461.
- Mendez, M. F., Mastri, A. R., Sung, J. H., and Frey, W. H. 1992. Clinically diagnosed Alzheimer disease: Neuropathologic findings in 650 cases. *Alzheimer Dis. Assoc. Disord.* 6: 35–43.
- Messa, C., Perani, D., Lucignani, G., Zenorini, A., Zito, F., Rizzo, G., Grassi, F., Del Sole, A., Franceschi, M., Gilardi, M. C., *et al.* 1994. High-resolution technetium-99m-HMPAO SPECT in patients with probable Alzheimer's disease: Comparison with fluorine-18-FDG PET. *J. Nucl. Med.* **35**: 210–216.
- Metz, C. E. 1978. Basic principles of ROC analysis. Semin. Nuclear Med. 8: 283–298.
- Mielke, R., Herholz, K., Grond, M., Kessler, J., and Heiss, W. D. 1992. Differences of regional cerebral glucose metabolism between presenile and senile dementia of Alzheimer type. *Neurobiol. Aging* 13: 93–98.
- Mielke, R., Pietrzyk, U., Jacobs, A., Fink, G. R., Ichimiya, A., Kessler, J., Herholz, K., and Heiss, W. D. 1994. HMPAO SPET and FDG PET in Alzheimer's disease and vascular dementia: comparison of perfusion and metabolic pattern. *Eur. J. Nucl. Med.* 21: 1052–1060.
- Minoshima, S., Frey, K. A., Koeppe, R. A., Foster, N. L., and Kuhl, D. E. 1995. A diagnostic approach in Alzheimer's disease using three-dimensional stereotactic surface projections of fluorine-18-FDG PET. J. Nucl. Med. 36: 1238–1248.
- Morris, J. C. 1993. The Clinical Dementia Rating (CDR): Current version and scoring rules [see comments]. *Neurology* **43**: 2412–2414.
- Morris, J. C., McKeel, D. W., Jr., Fulling, K., Torack, R. M., and Berg, L. 1988. Validation of clinical diagnostic criteria for Alzheimer's disease. Ann. Neurol. 24: 17–22.
- Mungas, D., Marshall, S. C., Weldon, M., Haan, M., and Reed, B. R. 1996. Age and education correction of Mini-Mental State Examination for English and Spanish-speaking elderly. *Neurology* **46**: 700–706.
- Perani, D., and Cappa, S. F. 2001. Brain imaging in normal aging and dementia. In Normal Aging and Dementia. Handbook of Neuropsychology (J. Grafman and F. Boller, Eds.). Elsevier, Amsterdam.
- Petit-Taboue, M. C., Landeau, B., Desson, J. F., Desgranges, B., and Baron, J. C. 1998. Effects of healthy aging on the regional cerebral metabolic rate of glucose assessed with statistical parametric mapping. *NeuroImage* 7: 176–184.
- Pietrzyk, U., Herholz, K., Schuster, A., von Stockhausen, H. M., Lucht, H., and Heiss, W. D. 1996. Clinical applications of registra-

tion and fusion of multimodality brain images from PET, SPECT, CT, and MRI. *Eur. J. Radiol.* **21:** 174–182.

- Rapoport, S. I. 1991. Positron emission tomography in Alzheimer's disease in relation to disease pathogenesis—A critical review. *Cerebrovasc. Brain Metab. Rev.* **3**: 297–335.
- Rapoport, S. I. 1997. Discriminant analysis of brain imaging data identifies subjects with early Alzheimer's disease. *Int. Psychogeriatr.* **9** (Suppl. 1): 229–235.
- Reiman, E. M., Caselli, R. J., Yun, L. S., Chen, K., Bandy, D., Minoshima, S., Thibodeau, S. N., and Osborne, D. 1996. Preclinical evidence of Alzheimer's disease in persons homozygous for the epsilon 4 allele for apolipoprotein E. N. Engl. J. Med. 334: 752– 758.
- Reisberg, B., Ferris, S. H., DeLeon, M. J., and Crook, T. 1982. The global deterioration scale for assessment of primary degenerative dementia. *Am. J. Psych.* **139**: 1136–1139.
- Salmon, E., Collette, F., Degueldre, C., Lemaire, C., and Franck, G. 2000. Voxel-based analysis of confounding effects of age and dementia severity on cerebral metabolism in Alzheimer's disease. *Human Brain Mapping* 10: 39–48.
- Salmon, E., Sadzot, B., Maquet, P., Degueldre, C., Lemaire, C., Rigo, P., Comar, D., and Franck, G. 1994. Differential diagnosis of Alzheimer's disease with PET. J. Nucl. Med. 35: 391–398.
- Signorini, M., Paulesu, E., Friston, K., Perani, D., Colleluori, A., Lucignani, G., Grassi, F., Bettinardi, V., Frackowiak, R. S. J., and Fazio, F. 1999. Rapid assessment of regional cerebral metabolic abnormalities in single subjects with quantitative and nonquanti-

tative [18F]FDG PET: A clinical validation of statistical parametric mapping. *NeuroImage* **9:** 63–80.

- Silverman, D. H., Small, G. W., Chang, C. Y., Lu, C. S., Kung de Aburto, M. A., Chen, W., Czemin, J., Rapoport, S. I., Pietrini, P., Alexander, G. E., Schapiro, M. B., Jagust, W. J., Hoffman, J. M., Welsh-Bohmer, K. A., Alavi, A., Clark, C. M., Salmon, E., de Leon, M. J., Mielke, R., Cummings, J. L., Kowell, A. P., Gambhir, S. S., Hoh, C. K., and Phelps, M. E. 2001. Positron emission tomography in evaluation of dementia: Regional brain metabolism and longterm outcome. JAMA 286: 2120–2127.
- Small, G. W., Ercoli, L. M., Silverman, D. H., Huang, S. C., Komo, S., Bookheimer, S. Y., Lavretsky, H., Miller, K., Siddarth, P., Rasgon, N. L., Mazziotta, J. C., Saxena, S., Wu, H. M., Mega, M. S., Cummings, J. L., Saunders, A. M., Pericak-Vance, M. A., Roses, A. D., Barrio, J. R., and Phelps, M. E. 2000. Cerebral metabolic and cognitive decline in persons at genetic risk for Alzheimer's disease. *Proc. Natl. Acad. Sci. USA* **97**: 6037–6042.
- Steel, R. D. G., and Torrie, J. H. 1980 *Principles and Procedures of Statistics. A Biometrical Approach.* McGraw-Hill, New York.
- Thompson, P. M., Moussai, J., Zohoori, S., Goldkorn, A., Khan, A. A., Mega, M. S., Small, G. W., Cummings, J. L., and Toga, A. W. 1998. Cortical variability and asymmetry in normal aging and Alzheimer's disease. *Cereb. Cortex* 8: 492–509.
- Thompson, P. M., Schwartz, C., Lin, R. T., Khan, A. A., and Toga, A. W. 1996. Three-dimensional statistical analysis of sulcal variability in the human brain. *J. Neurosci.* **16**: 4261–4274.
- Wechsler, D. 1987. Wechsler Memory Scale-Revised. Psychological Corp., New York.